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RESEARCHARTICLE

SOME MODIFIED EXPONENTIAL TYPE UNBIASED ESTIMATORS USING AUXILIARY ATTRIBUTE  
IN SIMPLE RANDOM SAMPLING

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ABSTRACT

The main concentration of this paper is to suggest modified exponential type unbiased ratio and product estimators using auxiliary attribute in simple random sampling. A comparative analysis of efficiency is carried out between suggested and existing estimators theoretically as well as numerically.

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INTRODUCTION

The estimation of the population parameters is a major challenge in sampling theory and many researchers in this field make serious efforts in the direction of efficiency and precision improvement of estimators of unknown population parameter of interest when study variable is highly correlated with the auxiliary variable. When the correlation between study variable and auxiliary variable is positive we use the ratio estimator and if correlation is negative we use the product estimator. In many situations auxiliary information is ignored when information is qualitative in nature such as sex and height of the person, amount of milk produced and a particular breed of the cow, amount of yield of wheat crop and a particular variety of wheat etc. (Shabbir and Gupta, 2010). In such situations, taking the advantage of point bi-serial correlation between the study variable y and the auxiliary attributes the estimators of population parameter of interest can be constructed by using prior knowledge of the population parameter of auxiliary attribute. The concentration of this paper is to suggest modified exponential type unbiased ratio and product estimators for estimating the population mean using auxiliary attribute in simple random sampling. The unbiasedness and variance of the suggested estimators have been obtained. A comparative analysis of efficiency is carried out between suggested and existing estimators theoretically as well as numerically.

Suggested Sampling Scheme and Notation

By considering a finite population which has N identifiable units  $\Theta_i (1 \leq i \leq N)$  and using simple random sampling without replacement (SRSWOR) a sample of size n is drawn from a population of size N. Here  $\Phi_i$  and  $y_i$  denote the observations on the variable  $\Phi$  and y respectively for  $i^{th}$  unit ( $i=1, 2, \dots, N$ ). Suppose there is a complete dichotomy in the population with respect to the presence or absence of an attribute, say  $\Phi$ , and it is assumed that attribute  $\Phi$  takes only two values 0 and 1 according as

$$\Phi_i = 1 \text{ if } i^{th} \text{ unit of the population possesses attribute } \Phi \\ = 0 \text{ otherwise.}$$

Here  $P = \frac{\sum_{i=1}^N \Phi_i}{N}$  and  $p = \frac{\sum_{i=1}^n \Phi_i}{n}$  denote the proportion of units in the population and sample respectively possessing attribute  $\Phi$ . Let  $\bar{y} = \frac{1}{n} \sum_{i=1}^n y_i$  and  $p = \frac{1}{n} \sum_{i=1}^n \Phi_i$  be the sample means of variable of interest y and auxiliary attribute  $\Phi$ ,  $\bar{Y} = \frac{1}{N} \sum_{i=1}^N y_i$  and  $P = \frac{1}{N} \sum_{i=1}^N \Phi_i$  be the corresponding population means. Let  $s_y^2 = \frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y})^2$  and  $s_\Phi^2 = \frac{1}{n-1} \sum_{i=1}^n (\Phi_i - p)^2$  be the

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sample variance and  $S_y^2 = \frac{1}{N-1} \sum_{i=1}^N (y_i - \bar{Y})^2$  and

$S_\Phi^2 = \frac{1}{N-1} \sum_{i=1}^N (\Phi_i - P)^2$  be the corresponding population

variance. Let  $C_y = \frac{S_y}{\bar{Y}}$  and  $C_\Phi = \frac{S_\Phi}{P}$ . Let  $\rho_{y,\Phi} = \frac{S_{y,\Phi}}{S_y S_\Phi}$  be the

point bi-serial correlation coefficient between  $y$  and  $\Phi$ . To determine the characteristic of the suggested estimators and existing estimators considered here, we define the following

terms,  $Y_y = \frac{\bar{y} - \bar{Y}}{\bar{Y}}$  and  $Y_\Phi = \frac{p - P}{P}$  such that

$$E[Y_i] = 0 \text{ for } (i = y, \Phi), E(Y_y^2) = \Psi \frac{S_y^2}{\bar{Y}^2},$$

$$E(Y_\Phi^2) = \Psi \frac{S_\Phi^2}{P^2} \text{ and } E(Y_y Y_\Phi) = \Psi \rho_{y,\Phi} \frac{S_y}{\bar{Y}} \frac{S_\Phi}{P} \text{ where}$$

$$\Psi = \left( \frac{1}{n} - \frac{1}{N} \right).$$

**Existing Estimators**

In this section, we consider the several existing estimators which are used for the estimation of population mean.

**Table 1. Existing ratio and product estimators using auxiliary attribute in simple random sampling.**

| S.No. | Estimators  | Mean Square Error [MSE(■)]  |
|-------|---|---|
| 1.    | $\theta_1 = \bar{y} \exp\left(\frac{P-p}{P+p}\right)$<br>[Bahl&Tuteja (1991) - Ratio Estimator] | $MSE(\theta_1) = \Psi \bar{Y}^2 \left[ C_y^2 + \frac{1}{4} C_\Phi^2 - \rho_{y,\Phi} C_y C_\Phi \right]$ |
| 2.    | $\theta_2 = \bar{y} \exp\left(\frac{p-P}{p+P}\right)$<br>[Bahl&Tuteja (1991)-Product Estimator] | $MSE(\theta_2) = \Psi \bar{Y}^2 \left[ C_y^2 + \frac{1}{4} C_\Phi^2 + \rho_{y,\Phi} C_y C_\Phi \right]$ |
| 3.    | $\theta_3 = \left(\frac{\bar{y}}{p}\right)P$<br>[Naik& Gupta (1996) - Ratio Estimator]          | $MSE(\theta_3) = \Psi \left[ S_y^2 + R^2 S_\Phi^2 - 2R\rho_{y,\Phi} S_y S_\Phi \right]$                 |
| 4.    | $\theta_4 = \left(\frac{\bar{y}}{P}\right)p$<br>[Naik& Gupta (1996) -Product Estimator]         | $MSE(\theta_4) = \Psi \left[ S_y^2 + R^2 S_\Phi^2 - 2R\rho_{y,\Phi} S_y S_\Phi \right]$                 |

**Table 2. Values of Parameters**

|          |                  |             |                         |          |               |                 |
|----------|------------------|-------------|-------------------------|----------|---------------|-----------------|
| $N = 89$ | $\bar{Y} = 3.36$ | $P = 0.124$ | $\rho_{y,\Phi} = 0.766$ | $n = 23$ | $C_y = 0.604$ | $C_\Phi = 2.19$ |
|----------|------------------|-------------|-------------------------|----------|---------------|-----------------|

**Table 3. MSEs and Percentage relative efficiency**

| Estimator  | MSEs   | Percent Relative Efficiency |                   |                   |                   |
|------------|--------|-----------------------------|-------------------|-------------------|-------------------|
|            |        | w.r.t $\theta_1$            | w.r.t. $\theta_2$ | w.r.t. $\theta_3$ | w.r.t. $\theta_4$ |
| $\theta_1$ | 0.2841 | 100                         | ■                 | ■                 | ■                 |
| $\theta_2$ | 0.8696 | ■                           | 100               | ■                 | ■                 |
| $\theta_3$ | 1.0836 | ■                           | ■                 | 100               | ■                 |
| $\theta_4$ | 1.9511 | ■                           | ■                 | ■                 | 100               |
| $\theta_5$ | 0.1960 | 144                         | 443               | 552               | 995               |
| $\theta_6$ | 0.2132 | 133                         | 407               | 508               | 915               |

**Suggested Estimators and its Properties**

In this section, we propose the modified exponential type ratio and product estimators by using sampling design defined in section 2 for estimation of population mean as:

Modified exponential type ratio estimator

$$\theta_5 = \left[ \bar{y} - (e^{A_1} - 1) \right] \dots \dots \dots (1)$$

Modified exponential type product estimator

$$\theta_6 = \left[ \bar{y} - (e^{A_2} - 1) \right] \dots \dots \dots (2)$$

where  $A_1 = \left[ P - \frac{NP - np}{N - n} \right]$  and  $A_2 = \left[ \frac{NP - np}{N - n} - P \right]$

To obtain the unbiasedness and variance of  $\theta_5$  and  $\theta_6$  up to the first order of approximation, equation (1) and (2) we expand in terms of  $Y$ 's .

The right hand side of equation (1) expand up to the first order of approximation in terms of  $Y$ 's

Rewriting  $\theta_5$  as

$$\theta_5 = \bar{Y} [1 + \Upsilon_y] - \left[ \frac{n P \Upsilon_\Phi}{N - n} \right] \dots \dots \dots (3)$$

it is clear that  $E(\theta_5) = \bar{Y}$ , after taking expectation on both sides of equation (3), i.e. up to the first order of approximation

$\theta_5$  is an unbiased estimator of the population mean  $\bar{Y}$ .

Now variance of  $\theta_5$  can be calculated as

$$Var(\theta_5) = E[\theta_5 - E(\theta_5)]^2$$

After neglecting the higher order terms

$$Var(\theta_5) = E \left[ \bar{Y} \Upsilon_y - \frac{n}{N-n} P \Upsilon_\Phi \right]$$

$$Var(\theta_5)$$

$$= \bar{Y}^2 E(\Upsilon_y^2) + P^2 \left( \frac{n}{N-n} \right)^2 E(\Upsilon_\Phi^2) - \frac{2n}{N-n} \bar{Y} P E(\Upsilon_y \Upsilon_\Phi)$$

After simplification variance of  $\theta_5$  is

$$Var(\theta_5) = \Psi \left[ S_y^2 + \left( \frac{n}{N-n} \right)^2 S_\Phi^2 - \frac{2n}{N-n} \rho_{y\Phi} S_y S_\Phi \right] \dots \dots \dots (4)$$

Same as , right hand side of equation (2), expand up to the first order of approximation in terms of  $\Upsilon' S$

$$\theta_6 = \bar{Y} [1 + \Upsilon_y] - \left[ \left\{ 1 + \frac{NP - nP(1 + \Upsilon_\Phi)}{N - n} - P \right\} - 1 \right]$$

Rewriting  $\theta_6$  as

$$\theta_6 = \bar{Y} [1 + \Upsilon_y] + \left[ \frac{n P \Upsilon_\Phi}{N - n} \right] \dots \dots \dots (5)$$

We can easily prove that  $E(\theta_6) = \bar{Y}$  i.e.  $\theta_6$  is an unbiased estimator of the population mean  $\bar{Y}$ .

The variance of  $\theta_6$  can be calculated as

$$Var(\theta_6) = E[\theta_6 - E(\theta_6)]^2$$

$$= E \left[ \bar{Y} \Upsilon_y + \frac{n}{N-n} P \Upsilon_\Phi \right] \text{ (neglecting the higher order terms)}$$

$$= \bar{Y}^2 E(\Upsilon_y^2) + P^2 \left( \frac{n}{N-n} \right)^2 E(\Upsilon_\Phi^2) + \frac{2n}{N-n} \bar{Y} P E(\Upsilon_y \Upsilon_\Phi)$$

After simplification variance of  $\theta_6$  is

$$Var(\theta_6) = \Psi \left[ S_y^2 + \left( \frac{n}{N-n} \right)^2 S_\Phi^2 + \frac{2n}{N-n} \rho_{y\Phi} S_y S_\Phi \right] \dots \dots \dots (6)$$

**Efficiency Comparison**

In this section we compare efficiency of the suggested estimators with existing estimators defined in section 2 and we derive the conditions in which suggested estimators are better than the existing estimators:

**Observation(1)**

$$MSE(\theta_1) - Var(\theta_5) = [-\rho_{y\Phi} R S_y S_\Phi + \frac{1}{4} R^2 S_\Phi^2 - \left( \frac{n}{N-n} \right)^2 S_\Phi^2 + \frac{2n}{N-n} \rho_{y\Phi} S_y S_\Phi] > 0$$

$$\text{if } \rho_{y\Phi} > \frac{S_\Phi}{S_y} \left[ \left( \frac{n}{N-n} \right)^2 - \frac{R^2}{4} \right] / \left[ \frac{2n}{N-n} - R \right]$$

**Observation (2)**

$$MSE(\theta_2) - Var(\theta_6) = [-\rho_{y\Phi} R S_y S_\Phi + \frac{1}{4} R^2 S_\Phi^2 - \left( \frac{n}{N-n} \right)^2 S_\Phi^2 - \frac{2n}{N-n} \rho_{y\Phi} S_y S_\Phi] > 0$$

$$\text{if } \rho_{y\Phi} < -\frac{S_\Phi}{S_y} \left[ \left( \frac{n}{N-n} \right)^2 - \frac{R^2}{4} \right] / \left[ \frac{2n}{N-n} - R \right]$$

**Observation (3)**

$$MSE(\theta_3) - Var(\theta_5) = [R^2 S_\Phi^2 - 2R \rho_{y\Phi} S_y S_\Phi - \left( \frac{n}{N-n} \right)^2 S_\Phi^2 + \frac{2n}{N-n} \rho_{y\Phi} S_y S_\Phi] > 0$$

$$\text{if } \rho_{y\Phi} > \frac{S_\Phi}{2S_y} \left( \frac{n}{N-n} + R \right)$$

**Observation(4)**

$$MSE(\theta_4) - Var(\theta_6) = [R^2 S_\Phi^2 - 2R \rho_{y\Phi} S_y S_\Phi - \left( \frac{n}{N-n} \right)^2 S_\Phi^2 - \frac{2n}{N-n} \rho_{y\Phi} S_y S_\Phi] > 0$$

$$\text{if } \rho_{y\Phi} < -\frac{S_\Phi}{2S_y} \left( \frac{n}{N-n} + R \right)$$

**Numerical Study**

Now we compare the performance of various estimators considered here using the data sets as previously used by Shabbir and Gupta (2010).

**Population:** (Source: Sukhatme and Sukhatme (1970), pp. 256).

$y$  = Number of villages in the circles.

$\Phi$  = A circle consisting more than five villages.

**Conclusion**

In section 4, we suggested modified exponential type unbiased ratio and product estimators and also obtained the characteristic of the suggested estimators. In section 5, we derive the general conditions in which we can say the suggested estimators are always more efficient than the existing estimators. In section 6, numerical study is done using the data used by Shabbir and

Gupta (2010). In numerical study, we see that the suggested estimators have highest percent relative efficiency w.r.t to all existing estimators defined in section 2.

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